

Isochores Exhibit Evidence of Genes Interacting With the Large-Scale Genomic Environment

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ABSTRACT

The genomes of mammals and birds can be partitioned into megabase-long regions, termed isochores, with consistently high, or low, average $C + G$ content. Isochores with high CG contain a mixture of CG-rich and AT-rich genes, while high-AT isochores contain predominantly AT-rich genes. The two gene populations in the high-CG isochores are functionally distinguishable by statistical analysis of their gene ontology categories. However, the aggregate of the two populations in CG isochores is not statistically distinct from AT-rich genes in AT isochores. Genes tend to be located at local extrema of composition within the isochores, indicating that the CG-enriching mechanism acted differently when near to genes. On the other hand, maximum-likelihood reconstruction of molecular phylogenetic trees shows that branch lengths (evolutionary distances) for third codon positions in CG-rich genes are not substantially larger than those for AT-rich genes. In the context of neutral mutation theory this argues against any strong positive selection. Disparate features of isochores might be explained by a model in which about half of all genes functionally require AT richness, while, in warm-blooded organisms, about half the genome (in large coherent blocks) acquired a strong bias for mutations to CG. Using mutations in CG-rich genes as convenient indicators, we show that $\approx 20\%$ of amino acids in proteins are broadly substitutable, without regard to chemical similarity.

ISOCORES, so named by Bernardi (BERNARDI *et al.* 1985; BERNARDI 2000), are large regions in the human genome, as long as tens of megabases, that are anomalously rich in C and G nucleotides. Isochores analogous to human are found in the genomes of all mammals and birds (BERNARDI 2000), plus a small number of additional reptiles such as the Nile crocodile (HUGHES *et al.* 1999). Invertebrates, and almost all cold-blooded vertebrates, do not manifest isochore structure in their genomes. The putative common ancestor in which isochores originated is thus an amniote in the Carboniferous period [≈ 300 million years before present (MYBP)], although it was not until after the Permian–Triassic extinction (≈ 250 MYBP) that the carriers of isochores, namely archosaurs, birds, and mammals, proliferated.

Isochores are by no means subtle features in the genome (IHGSC 2001). By way of example, Figure 1 shows the $A + T$ (complement of $C + G$) content of three human, and three zebrafish, chromosomes, plotted on a common scale. The nucleotide counts are shown as bars in 300-kb bins, with the base of the bars at $A + T = 0.58$, an arbitrary value that approximately divides CG isochores from AT isochores (as we refer to regions that are not CG rich).

It is not a settled issue whether isochore formation continues today, that is, whether CG isochores are continuing to form from AT isochores. However, a body of recent evidence suggests that, on the contrary, isochores are gradually disappearing from mammalian genomes (DURET *et al.* 2002; BELLE *et al.* 2004). If so, then we may view isochores as fossils of a unique period in our past during which a strong mutational pressure first appeared and then disappeared. Apart from the obvious question as to what caused this to happen, we may also hope to learn from the isochore-forming event something about the interaction of genes, as primary carriers of functional information, with the much larger genome that they inhabit.

It has proved surprisingly difficult to find functional relationships between isochores and the genes inside them (IHGSC 2001; VINOGRADOV 2003). By default, the more conservative view has been that isochores are predominantly the result of the accumulation of selection-neutral changes caused by (evidently spatially nonuniform) mutation or repair biases. One currently favored model is biased gene conversion (BGC) during homologous recombination (EYRE-WALKER and HURST 2001). If isochore evolution is selection neutral, then genes are passive riders on the isochore background. That is, their noncritical elements, such as synonymous bases in third codon positions and nonfunctional bases in their 3'- and 5'-untranslated regions (UTRs), should evolve toward CG richness along with the rest of an isochore.

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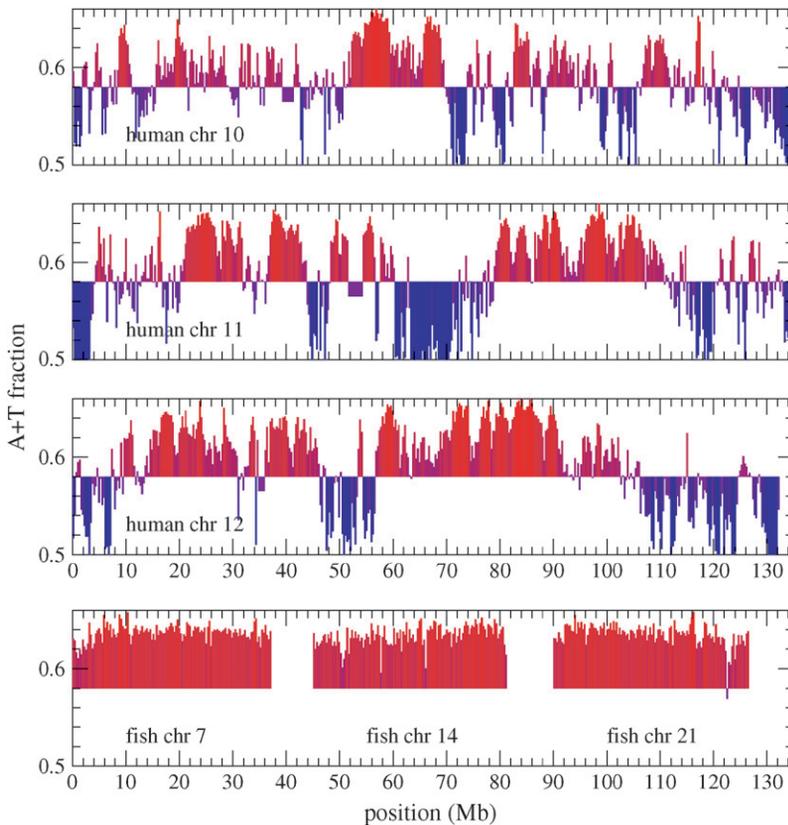


FIGURE 1.—Local $A + T$ fraction of typical human and zebrafish chromosomes. Counts are shown in nonoverlapping 300-kb windows.

Indeed, it is well established (BERNARDI *et al.* 1985; CLAY *et al.* 1996; HAMADA *et al.* 2003), and easy to show, that the CG content of third codon positions and 3'- and 5'-UTRs are all strongly correlated with the CG content of the flanking genomic region.

Less conservative, but also longstanding, is the hypothesis that the evolution of isochores was favored by positive natural selection, for example, selection in warm-blooded vertebrates for DNA that is stable at higher temperature (BERNARDI 2000; SMITH and EYRE-WALKER 2001). However, several such hypotheses notwithstanding, the nature of the selection pressure remains obscure (EYRE-WALKER and HURST 2001; BELLE *et al.* 2002).

If isochore formation was predominantly selection neutral, then there should not be statistically significant functional differences between genes in an AT *vs.* CG isochore, since during isochore formation the (pre-existing) population of genes is simply hitchhikers. However, without reference to isochores, we have previously shown (ROBINS and PRESS 2005) that AT-rich and CG-rich genes are readily distinguishable, statistically, by gene functionality. In particular, AT-rich genes are preferentially associated with one set of biological processes, centered on transcription and mRNA processing, while CG-rich genes are associated with another set, centered on signal transduction, receptors, and signaling cascades. Can this finding be reconciled with a selection-neutral model for isochores?

We will see below that the answer is yes, but with an important caveat. As one would expect, AT- and CG-rich genes are associated with the corresponding AT and CG isochores. But the association is not one-to-one: Genes in AT isochores are predominantly AT rich, while genes in CG isochores can be either AT rich or CG rich, resulting in a complex landscape of AT-rich intrusions into what are otherwise CG isochores. The two groups of genes in CG isochores, AT rich and CG rich, are in fact statistically distinguishable by function. However, when one aggregates both groups of genes in CG isochores, one obtains a mixture that is not functionally distinguishable from the genes in AT isochores, consistent with the neutral model.

Thus, while the correlation of function with CG richness in CG isochores is clear evidence of selection, it may indicate only negative selection. That is, the evolutionary pressure toward CG richness could be entirely the result of neutral mutations; but in that case some genes, correlated by function, felt negative fitness pressure to resist the neutral mutations and remained AT-rich genes within a CG isochore.

On the other hand, and arguing for isochore formation by positive selection, we find evidence that genes that became CG rich (in CG isochores) are far from passive passengers: They are *more* CG rich than their surroundings. That genes are at special locations of composition is already suggested visually, at least for CG-rich genes, if one simply looks at the position and

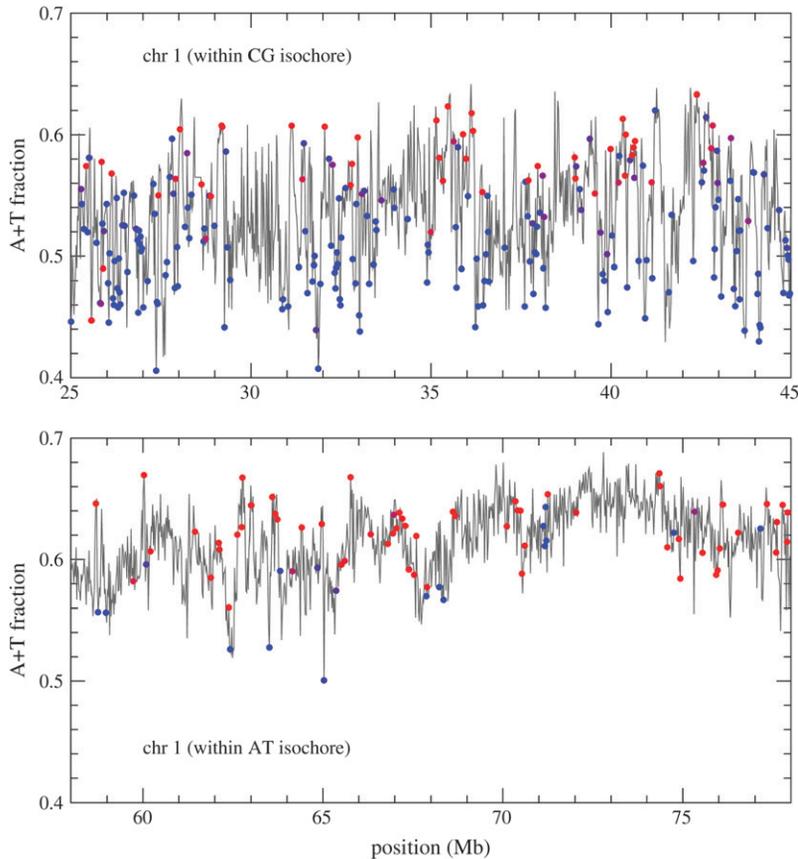


FIGURE 2.—Two regions of human chromosome 1, plotting $A + T$ counts in 20-kb windows, and showing the location of all RefSeq genes. Genes are plotted at the $A + T$ value of the window in which they occur, but with their color continuously varying from red (AT-rich gene) to blue (CG-rich gene). Genes in a CG isochore (top plot), notably CG-rich genes, tend to be more extreme than their flanking regions; there is less such tendency in an AT isochore (bottom plot).

composition of genes relative to window counts (Figure 2), where an unexpected number of CG-rich genes seem to occur in bins that are extrema. We give a more quantitative test below.

METHODS

Defining gene populations and large-scale isochores:

We use $A + T$ and $C - G$ counts in a gene's 3'-UTR to determine whether it is an AT-rich or a CG-rich gene, applying the algorithm given in ROBINS and PRESS (2005), Equation 1, to get a probability. This method was shown to yield the cleanest separation of the two gene populations.

Since isochores are not homogeneous (IHGSC 2001, and *cf.* Figure 1), a precise definition is perforce somewhat arbitrary. However, if one plots the above AT- *vs.* CG-rich probability for each gene along the genome, as in Figure 3, a clear pattern emerges: Some regions extending over many megabases contain predominantly AT-rich genes, while other regions contain a more equal mixture of AT- and CG-rich genes. There are few, if any, large regions containing predominantly CG-rich genes, which is consistent with previous evidence (PAVLICEK *et al.* 2002) that CG isochores have larger compositional variances than do AT isochores.

We can therefore define isochore boundaries by a Markov model that alternates between two states, AT

dominant and mixed. In the AT-dominant state, the respective probabilities of an AT-rich and CG-rich gene are taken as (0.9, 0.1), while in the mixed state they are taken as (0.5, 0.5). The state transition probability between any two consecutive genes is taken as 0.001 (that is, 0.999 chance of remaining in the same isochore state). We then use the standard forward-backward method to find the probability, at each gene, of its being in the AT-dominant state (which we now term an AT isochore) or the mixed state (which we call a CG isochore). We find that this classification is quite insensitive to varying all of the parameters above. In particular the transition probability can be varied over orders of magnitude, because the multiplicative probabilities of a relatively small number of genes can easily force a state transition, even if its *a priori* probability is unrealistically small. Results are shown in Figure 4.

Comparing Figures 1 and 4, one sees that the above Markov model largely captures one's visual impressions of large-scale structure, but now objectively (at least up to choice of model parameters). We can also validate the gene-based model by comparing it to a similar Markov model that uses raw 300-kb window counts instead of genes, shown as the red line in Figure 4. In this model, we assign a 300-kb window to the high state if its count of $A + T > 0.565$, a not untypical value in the isochore literature (PAVLICEK *et al.* 2002). An AT isochore is taken to have high or low windows with respective

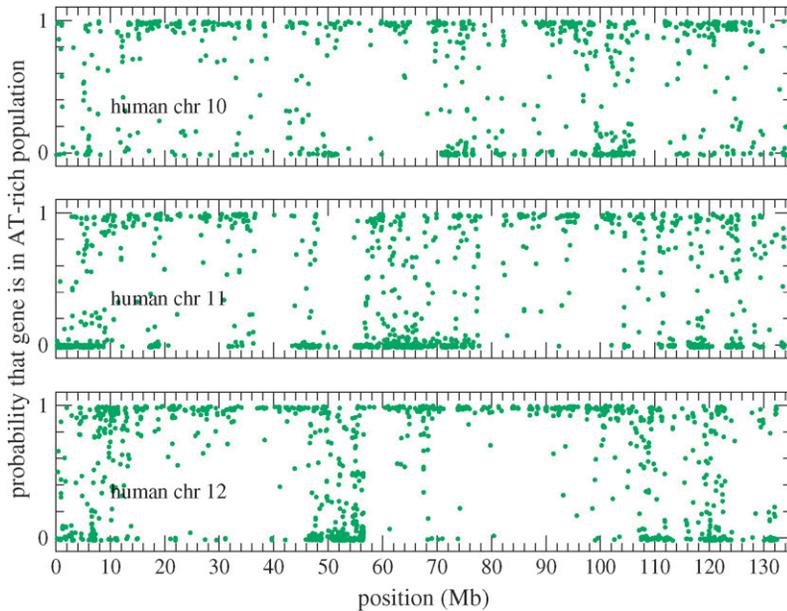


FIGURE 3.—RefSeq genes plotted according to their probability of being in the AT-rich population, for three typical chromosomes. Large regions of AT-rich genes, and of mixed AT- and CG-rich genes, are evident. Large regions of CG-rich genes are conspicuously absent.

probabilities (0.75, 0.25). A CG isochore has (0.5, 0.5), again reflecting its relatively larger variances. The transition probability is 0.001, as before. The results of this model are shown as the red line in Figure 4 and are insensitive to the adopted parameters. Our gene-based and window-based models for isochore identification agree in 93% of all locations in the human genome.

In characterizing variations on large, megabase scales, we necessarily miss smaller-scale features, predominantly AT-rich intrusions into CG isochores. These show up as an increase in the observed variance. It is a matter of semantics whether or not to regard these features as small isochores (COHEN *et al.* 2005).

Assessing two gene populations by GO score: In previous work (ROBINS and PRESS 2005) using gene ontology (GO) keyword counts, we characterized results by their statistical significance (t and P -values). Here, we want something more like a linear scale, so that a mixture of two populations will have a score that lies proportionally between the scores for the populations individually.

Using results from ROBINS and PRESS (2005), we define a set of “population-N” (for “nuclear”) indicator words as the following: nucleic acid, nucleus, transition metal, zinc, bound, ZNF*, RNA, mRNA, DNA, nucleobase, nucleoside, and translation. We define a set of “population-S” (for “signaling”) indicator words

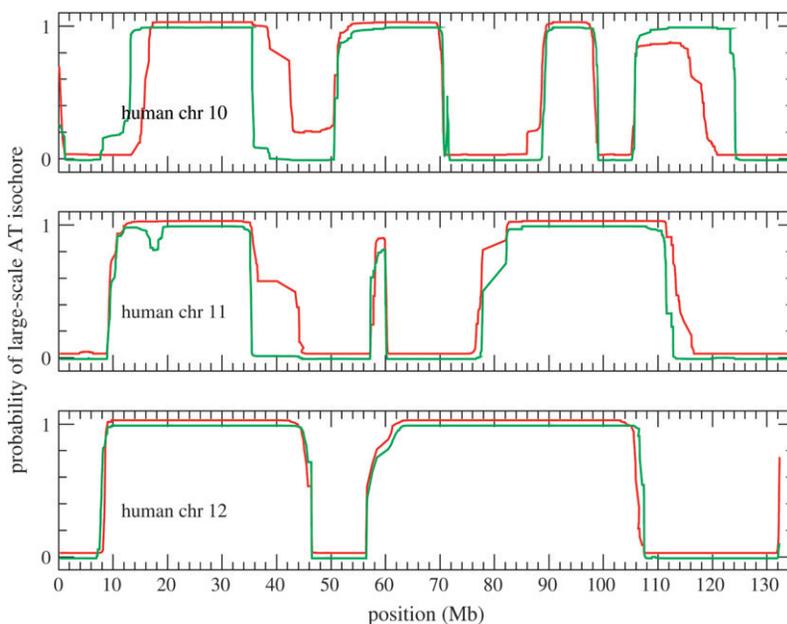


FIGURE 4.—Green line, isochore boundaries obtained by applying a Markov model with two states, “AT-rich genes” and “mixed genes.” Red line, isochore boundaries obtained by a similar model using raw counts in 300-kb windows.

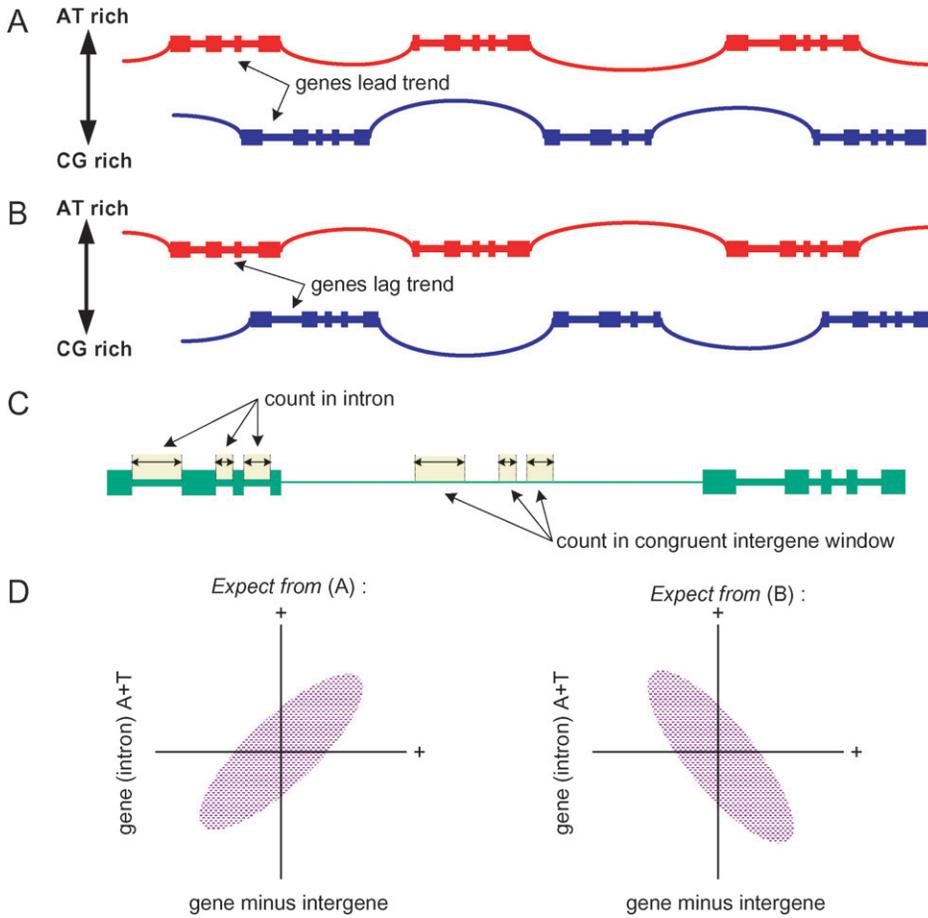


FIGURE 5.—Strategy for measuring whether genes are more (A) or less (B) compositionally extreme than their immediate surroundings. The gene’s composition is measured by counts in its introns only (C). Counts in the adjacent intergenic region are made with an identical window function. We expect (D) a positive correlation between gene and gene minus intergene if genes are compositionally more extreme and a negative correlation if less extreme.

as: signal transduction, signaling cascade, receptor, transducer, communication, signal, transmembrane, channel, immune, and pore. It is an important point that we did not choose these populations or words arbitrarily; rather, they emerged uniquely from the data as the word sets that most statistically significantly distinguish AT-rich and CG-rich genes (without regard to their locations in isochores).

Let N_N be the total number of occurrences of population-N words across the genome (e.g., in RefSeq genes) and N_S be the corresponding number for population-S words. Define $r_{NS} \equiv N_N/N_S$. (For the RefSeq genes we have $N_N = 31406$, $N_S = 16585$, and $r_{NS} = 1.89$.)

Now suppose that we have a large, probabilistically known, set of genes, meaning that we can assign a probability p_i of gene i ’s being in the set, and $\sum_i p_i \gg 10^2$ (say). Then we define that set’s “signaling minus nuclear score” (SMNS) by

$$\text{SMNS} \equiv \frac{r_{NS} \sum_i \sum_{j \in S} p_i \delta_{ij} - \sum_i \sum_{j \in N} p_i \delta_{ij}}{r_{NS} \sum_i \sum_{j \in S} p_i \delta_{ij} + \sum_i \sum_{j \in N} p_i \delta_{ij}} \quad (1)$$

Here S is the set of population-S words, N is the set of population-N words, and δ_{ij} is 1 if word j occurs for gene i and zero otherwise. By construction, SMNS of the

whole genome is zero. It is 1 for a set of genes that have no population-N words and -1 for a set of genes that have no population-S words.

Usefully, we can also estimate the error for the SMNS:

$$\sigma(\text{SMNS}) \approx \frac{\sqrt{r_{NS}^2 \sum_i \sum_{j \in S} p_i^2 \delta_{ij} + \sum_i \sum_{j \in N} p_i^2 \delta_{ij}}}{r_{NS} \sum_i \sum_{j \in S} p_i \delta_{ij} + \sum_i \sum_{j \in N} p_i \delta_{ij}} \quad (2)$$

The approximation made is to ignore the error in the denominator of Equation 1 as compared to that of the numerator. This is because (with foresight) it will turn out that the SMNS score is never larger than a few tenths.

Equation 2 allows us to compare different sets of genes for statistically significantly different SMNSs.

Determining whether genes are more or less compositionally extreme: As discussed above, it is important to have an objective measure of whether genes are more or less extreme in $C + G$ or $A + T$ than their immediate surroundings. One measure of this tendency is to compare $A + T$ at a gene’s location with $A + T$ at the midpoint of the intergenic region between the gene and its next neighbor. Referring to Figure 5, if genes are more compositionally extreme (as shown in Figure 5A) we should get a different correlation

between gene and intergene than if genes are less compositionally extreme (as shown in Figure 5B). Figure 5D shows the two cases schematically.

A difference between the variance of genes and that of intergenes due to any other effect can confound the proposed measurement. For example, if genes had a smaller variance in their $A + T$ composition for functional reasons, this would bias the measurement toward Figure 5B. Or, if the measurement accuracy of $A + T$ were poorer for genes (due to a smaller counting length) than for intergenes, then Figure 5A would be erroneously favored. To mitigate these kinds of systematic errors, we adopt the strategy shown in Figure 5C: We characterize a gene's $A + T$ exclusively by its introns, which should have the least functional constraints; and we make intergenic counts with exactly the same window pattern as that of the gene to which they are being compared. If there are residual systematic biases in the introns (which do contain some functionality), we expect them to show up as a systematic shift in $A + T$, not a change in the variance. (In fact, below, we see such small shifts.) The signature of genes that are compositionally more extreme than their surroundings is a positive correlation between gene and gene minus intergene. The signature of genes that are less compositionally extreme is a negative correlation.

Maximum-likelihood phylogeny and branch lengths:

Below, we construct phylogenetic trees by aligning orthologous genes in human, chicken (*Gallus gallus*), and frog (*Xenopus tropicalis*), with fish (*Danio rerio*) as an outgroup. We use only genes with orthologs in all four organisms, as reported by the Ensembl database (BIRNEY *et al.* 2006). (We have also checked that similar results are obtained if this constraint is relaxed.) We can construct independent trees for any particular population of genes, *e.g.*, AT rich or CG rich in CG isochores. In most cases (identified below) we use only fourfold degenerate third-codon positions, although, as we will see, interesting results are also obtained for nonsynonymous first and second codon positions.

The reconstruction method is the standard maximum-likelihood (ML) method (FELSENSTEIN 1981, 2004), based on a Markov evolutionary model along each branch. We assume the established tree topology among the four species. We allow completely general transition matrices [*e.g.*, not necessarily having the general time-reversible (GTR) form] and solve for a different transition matrix along each branch. This generality is possible because of the large amount of data available, yielding negligible statistical errors in the reconstruction. Errors are thus dominated by modeling errors, for example, violation of the Markov model assumption or non-i.i.d. of individual base positions; these modeling errors are, of course, difficult to assess quantitatively. The maximum-likelihood reconstruction is found iteratively by the EM method (DEMPSTER *et al.* 1977), alternating between the calculation of node probabilities separately for each

base position and the reestimation of the common (across base positions) set of transition matrices.

Because we do not assume time reversibility, the ML method is in principle capable of producing a rooted tree; that is, the “pulley principle” (FELSENSTEIN 1981) does not strictly apply. We find, however, that the location of the four-species common ancestor root is rather poorly determined by the data, indicating that deviations from time reversibility are small, at least along the path between the quadruped common ancestor and fish. We therefore use fish only as an outgroup and show, below, only the quadruped ancestor tree, which is accurately rooted (at least statistically).

Having obtained the transition (that is, base substitution) matrix \mathbf{A} for an edge, we resolve it into an infinitesimal generator matrix \mathbf{G} and a branch length μ , such that

$$\mathbf{A} = \exp(\mu\mathbf{G}) \quad (3)$$

with \mathbf{G} having zero row sums, zero or negative diagonal, and zero or positive off-diagonal elements. Since \mathbf{G} can absorb any constant factor from μ , it needs a normalization convention. A convenient one is

$$\text{tr}(\mathbf{G}) = -4. \quad (4)$$

Then μ is the evolutionary distance measured in mean changes per base for a (standardized) uniform nucleotide distribution, essentially equivalent to the standard log-det distance (LOCKHART *et al.* 1994; STEEL 1994) and closely related to the parilinear distance (LAKE 1994). (See GU and LI 1996 for a comparison of these distance measures.)

The generator matrix \mathbf{G} usefully encodes the mutational biases of individual mutation events. In the context of this article's interest in mutations from AT to CG (or vice versa), two useful summary values are the sums of all off-diagonal elements corresponding to transitions in one direction ($AT \rightarrow CG$) or the other ($CG \rightarrow AT$). Below, we refer to these values as “propensities” for each direction.

As a check on the ML reconstruction, we used all sets of pairwise (only) alignments among the four species. It is well known (CHANG 1996; LAKE 1997; BAAKE 1998) that the full transition matrices cannot be obtained from pairwise data alone. However, it is easy to get branch lengths from pairwise data. All the pairwise parilinear distances (which are additive both up and down the tree) give an overdetermined set of linear relations among the individual branch lengths. We solve for the best solution in the least-squares sense. Reassuringly, the lengths obtained by this method are almost identical to those obtained by ML reconstruction.

RESULTS

Characterizing the three gene groups: With the above methods, we can assign to each gene a probability of being in the AT-rich (*vs.* CG-rich) population and,

TABLE 1
RefSeq genes by gene and isochore AT or CG richness

Isochore type	Gene type		
	AT (%)	CG (%)	Total (%)
iAT	28	$\leq 7^a$	35
iCG	19	46	65
Total	47	53	100

^aiAT/CG genes are likely to be overcounted (see text).

separately, a probability of being in an AT (*vs.* CG) isochore. The results are shown in Table 1. We adopt the notations iAT and iCG as denoting isochores and AT and CG as denoting genes, so that the three principal populations are iAT/AT, iCG/CG, and iCG/AT. Although there are undoubtedly some genuine iAT/CG genes, many or most genes that we classify as iAT/CG are probably the result of misidentified isochore boundaries. Therefore we often restrict our attention to the three principal groups mentioned above. It is previously known that CG-rich regions have higher gene density and smaller gene lengths (IHGSC 2001).

The *A + T* fraction of genes classed as iAT/AT is significantly greater than that classed as iCG/AT, 53.6% *vs.* 46.0% (third codon position counts). Part of this difference is likely due to false positives from the larger number of iCG/CG genes, since the AT-rich and CG-rich gene populations are overlapping distributions. For iCG/CG genes, the *A + T* fraction is 30.1%.

GO signature is strong in CG isochores and weak or absent in AT isochores: The SMNS score was defined above to be zero over the average gene population, positive for groups of genes with population-S GO keywords (such as “signal transduction”), and negative for groups of genes with population-N GO keywords (such as “nucleic acid”). Scores, and uncertainties, for the four gene groups are as follows: 0.102 ± 0.006 for iCG/CG, -0.239 ± 0.009 for iCG/AT, -0.010 ± 0.009 for iAT/AT, and 0.019 ± 0.018 for iAT/CG (the larger uncertainty from the smaller population size).

What is remarkable is that the largest positive and negative scores, by far, are for genes in CG isochores, while genes in AT isochores have SMNS scores consistent with zero. In other words, AT-rich genes in CG isochores are functionally more extreme (population N) than AT-rich genes in AT isochores, even as their nearby neighbors on the genome, the iCG/CG genes, tend strongly to population-S functionality. This effect is not a correlation with AT richness—indeed, it has the opposite sign—since iCG/AT genes are markedly less AT rich than iAT/AT genes. The observed effect is likewise opposite to what would be expected from any misclassifying of iCG/CG genes as iCG/AT.

The average SMNS scores for genes in CG and AT isochores are, respectively, 0.003 ± 0.006 and -0.006 ± 0.008 , that is, statistically zero. It is striking that the CG

isochores are so accurately zero, since that value is obtained only by averaging a large positive (iCG/CG) and an even larger negative (iCG/AT) value in just the right proportions.

These data suggest that AT and CG isochores in fact contain the same mixtures of functionality (average SMNS zero), but that only in CG isochores have these differences been made visible as differences in gene AT richness. This is evidence that whatever “marked” large contiguous regions of the genome as incipient CG isochores did so without reference to the gene content within those regions. It is consistent with a scenario in which genes in AT isochores never experienced the pressure that created the isochores (whether neutral or selection), while genes in CG isochores were thus challenged, but with dramatically different (and functionally correlated) responses, varying from gene to gene.

Human genes are more compositionally extreme than their surroundings: Figure 6 shows the result of applying the methodology described above (and in Figure 5) to the human genome. A significant positive correlation between gene and gene-minus-intergene counts is seen for all three gene populations, most strongly for iCG/CG genes. This indicates that all genes have some tendency to be more extreme than their flanking sequence with respect to (depending on the gene) CG or AT richness. The tendency is by far strongest for CG-rich genes. As plotted, Figure 6 does not exclude repeating elements, but the results are not significantly different if we exclude either (i) all elements identified by RepeatMasker or (ii) only the most common LINE and SINE elements.

Phylogenetic reconstruction shows mutation bias but not strong positive selection: Figure 7 shows the result of constructing ML molecular phylogenies, separately for genes that are iAT/AT and iCG/CG in humans. The phylogenies are based on the alignment of orthologous genes, as described in METHODS, above. Figure 7 shows results for fourfold degenerate third codon positions. The red and blue arrows have areas proportional to the propensities for transitions in the direction AT \rightarrow CG (blue) or CG \rightarrow AT (red), as determined from the generator matrices **G** on each branch (see METHODS). The branch lengths are determined with a statistical accuracy of $\sim \Delta\mu \approx 0.008$ ($1-\sigma$), as determined by re-sampling, so all the differences shown in Figure 7 are highly statistically significant. (High statistical accuracy is obtainable because the amount of data is huge.)

In the human iAT/AT genes, one sees a high degree of consistency on all branches. The greater propensity toward AT results in an AT-rich equilibrium for the genes in all three species.

In the human iCG/CG genes, one sees for the branch between the common ancestor and frog about this same balance of propensities. For human and (to a lesser extent) chicken, however, one sees a strong mutational bias toward CG. Such a bias is not unexpected since we

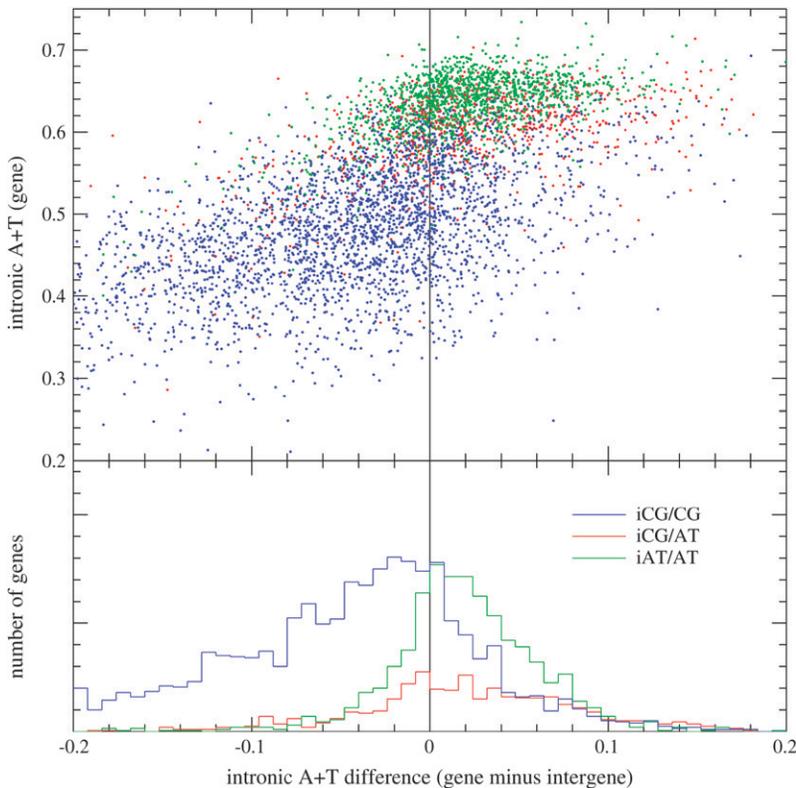


FIGURE 6.—Results of testing whether genes are more or less compositionally extreme than their surroundings. Blue, red, and green denote, respectively, CG genes in CG isochores (iCG/CG), AT genes in CG isochores (iCG/AT), and AT genes in AT isochores (iAT/AT). All three gene types tend to be more extreme than their immediate surroundings, most strongly for iCG/CG (compare Figure 5D).

have, of course, selected this sample for CG richness—and it had to come from somewhere.

What is most interesting in Figure 7 is what is *not* seen, namely any large disproportionate elongation for iCG/CG of the human and chicken branches relative to frog. Third codon positions are generally accepted as being governed dominantly by the neutral theory of molecular evolution (KIMURA 1983), albeit balanced by a moderate positive selection favoring the “major codons” for each amino acid (AKASHI 1994, 1996). Under the neutral theory, any branches on which positive isochore selection operates should be lengthened by a factor $\max(N_e s, 1)$, where N_e is the effective population size

and s is the positive selective advantage of a mutation per generation. On the other hand, there is no particular reason to think that $N_e s$ due to codon usage effects, and their effects on translation rates (LEVY *et al.* 1996; ZOLOTUKHIN *et al.* 1996; WELLS *et al.* 1999), should be very different on different branches of the quadruped ancestor tree.

While a small elongation of the human and chicken branches, on the order of ≈ 1.3 , may be present in the data, it would require a remarkable numerical coincidence among unrelated quantities, namely $s \approx 1/N_e$, to interpret this as positive selection. Rather, barring subtle competing effects, one might reasonably have expected

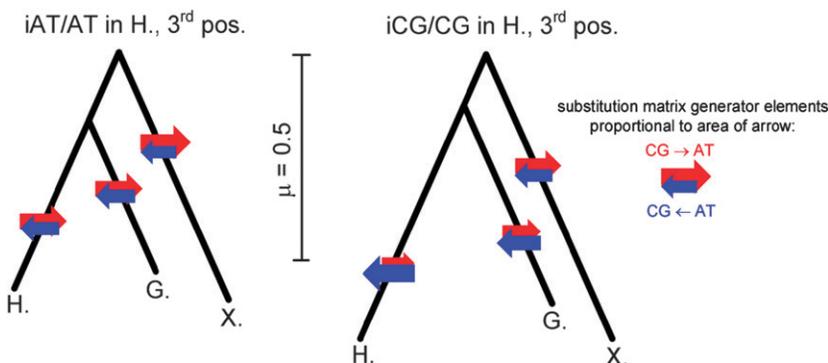


FIGURE 7.—Maximum-likelihood reconstruction of branch lengths and transition matrices derived from synonymous third codon positions for orthologous genes that are iAT/AT (left tree) and iCG/CG (right tree) in human. Shown are ancestral branches of human, chicken (*G. gallus*), and frog (*X. tropicalis*). Red and blue arrows summarize mutational biases toward AT or CG richness, respectively. The expected bias toward CG is seen in iCG/CG genes on the human and chicken branches. Not seen, however, is any large increase in apparent mutation rate (branch length), as would result from significant positive selection. Branch lengths μ are in units of mutations per base position. The trees are rooted by using an additional species (*D. rerio*) as an outgroup.

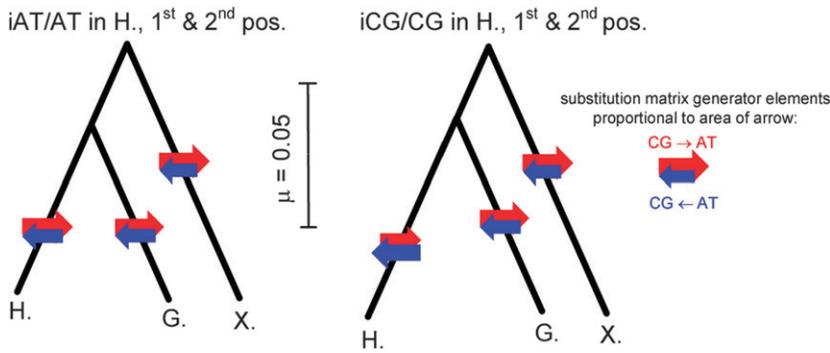


FIGURE 8.—The same as Figure 7, but now for first and second codon positions, where mutations cause protein amino acid changes. Note change of scale from Figure 7. Evolutionary distances are found to be reduced by a factor of ≈ 6 , but the figure is otherwise almost a scaled copy of Figure 7.

positive selection to manifest itself as a lengthening of, say, one or more powers of 10.

We can do the same molecular phylogenetic reconstruction on first and second codon positions, where mutations will (in general) result in protein amino acid changes, and which should therefore be functionally conserved. Results are shown in Figure 8. For both iCG/CG and iAT/AT genes, the trees for first and second codon positions are nearly identical to the trees for third codon positions, but scaled by a factor $f \approx \frac{1}{6}$, which (in Kimura's language) we can identify as the functional constraint, that is, the fraction of mutations that are approximately neutral. In Figure 8, the branch lengths are determined with a statistical accuracy of $\sim \Delta\mu \approx 0.0003$ ($1-\sigma$), so all the differences shown in Figure 8 are again highly statistically significant.

Since functional selection on first and second codon positions occurs at the protein level, quite different from functional selection on third codon positions due to codon usage bias, the consistency, up to a factor f , between Figures 7 and 8 is reassuring. It argues that the signature of positive selection for mutations to CG in iCG/CG genes is not being confounded by other functional effects. Such a signature, at least of any significant magnitude $N_e s \gg 1$, is simply not there.

Many protein amino acid changes are remarkably neutral: If isochore formation is indeed dominated by neutral mutation, as Figures 7 and 8 suggest, then isochores, and iCG/CG genes in particular, provide an interesting window into the question of the neutrality (or lack thereof) of amino acid-changing mutations. The leaf-to-leaf phylogenetic distances shown in Figure 8 for orthologous genes, functional in all the species compared, immediately show that $\geq 20\%$ of all amino acids can be mutated. However, by itself, this does not exclude such possibilities as (i) the mutations are under positive selection and reflect divergences in gene function or (ii) the mutations, while neutral, are allowed only between chemically similar amino acids.

The formation of iCG isochores in effect "labels" a set of mutations, identifiable at least statistically in iCG/CG genes, whose origin is unrelated to the function of any particular gene. We can then look at patterns of amino acid substitution across the reconstructed molecular

phylogeny. Particularly interesting are substitutions that correspond to *net* changes in amino acid usage, because these indicate broad trends, not gene-specific optimizations. We have done this comparison between human and fish. *D. rerio* was analyzed simultaneously with the species shown in Figures 7 and 8. Although, as an outgroup, it cannot be rooted, it is available for pairwise comparison between leaf (extant) taxa. There is of course no imputed direction of time in this pairwise comparison.

We have examined the aligned sequences of all human iCG/CG genes and their known zebrafish orthologs and counted the frequency with which amino acids are substituted. The resulting 20×20 table of counts may be looked at from two viewpoints: From a biochemical perspective, we may ask whether the substitution patterns "make sense" in favoring substitutions that are close in chemical property; or, from a genomic perspective, we may ask whether the substitutions seem driven by an exogenous pressure to increase $C + G$.

A first observation is that amino acid usage differs very significantly between human and fish coding regions. The difference is greatest between iCG/CG human genes and their fish orthologs and less for iAT/AT and iCG/AT. For example (Table 2), for iCG/CG genes, proline, alanine, and glycine usages are, respectively, 20, 19, and 13% higher in human than in fish, while asparagine, isoleucine, and methionine usages are, respectively, 21, 18, and 17% lower. One notes immediately that the former have exclusively C or G in the first and second codon positions, while the latter have A and T.

Figure 9 shows a biochemical perspective. For each of the eight amino acids with the greatest net positive or negative changes in usage between fish and human, arrows are shown indicating the four most frequent substitutions. The underlying diagram, after BETTS and RUSSELL (2003), puts closely substitutable amino acids close to each other. One sees that only a few of the most frequent observed substitutions make biochemical sense, *e.g.*, Lys (K) to Arg (R), or Ile (I) to Leu (L), while many others are, figuratively and literally, a stretch: Lys (K) to Ser (S), Leu (L) to Arg (R), Leu (L) to Pro (P), Thr (T) to Pro (P), etc. It is hard to imagine that there would not be significant functional consequences in making these kinds of substitutions in

TABLE 2
Amino acid usage changes from fish orthologs to human CG-rich genes

Amino acid	% change	Relevant codons	Largest four "came from" (when +%), or "went to" (when -%) and codon change			
Pro	+20	ccn	Ser (t → c)cn	Leu c(t → c)n	Ala (g → c)cn	Thr (a → c)cn
Ala	+19	gcn	Ser (t → g)cn	Thr (a → g)cn	Val g(t → c)n	Gly g(g → c)n
Gly	+13	ggn	Ser (a → g)gn	Ala g(c → g)n	Glu g(a → g)n	Arg (a → g)gn
Arg	+6	cg agn	Lys a(a → g)n	Gln c(a → g)n	Ser (tc → cg)n	Leu c(t → g)n
Lys	-13	aan	Arg a(a → g)n	Gln (a → c)an	Glu (a → g)an	Ser <i>complex</i>
Met	-17	atg	Leu (a → c)tg	Val (a → g)tg	Ile <i>at(g → n)</i>	Thr a(t → c)g
Ile	-18	atn	Val (a → g)tn	Leu (a → c)tn	Thr a(t → c)n	Ala (at → gc)n
Asn	-21	aan	Ser a(a → g)n	Asp (a → g)an	Gly (aa → gg)n	Thr a(a → c)n

All codon changes increase C + G except five in italics, which are neutral.

≈20% of particular amino acids, unless something like this fraction of amino acid positions in proteins is close to universally substitutable.

If the data in Figure 9 do not make sense biochemically, they do make sense when mapped into the genetic

code. As shown in Table 2, of the 32 largest substitutions, 27 can be explained as single mutations in the first or second codon that change an A or T into a C or G. The remaining 5 are all CG neutral. None of 32 are codon changes favoring A or T.

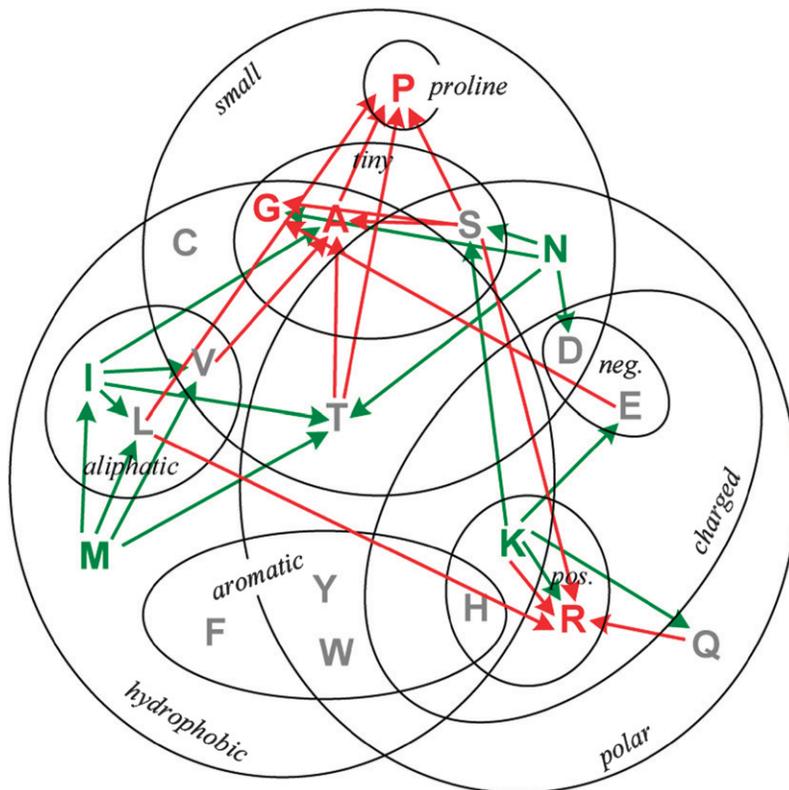


FIGURE 9.—Principal amino acid usage differences between fish and human orthologs, for CG-rich human genes in CG isochores. Green denotes fractionally most decreasing, and red most increasing, amino acids. For these, the four most frequent substitutions are shown. The direction of all arrows is a comparison from fish to human. (This is a graphical convention, not an arrow of time.) The observed trends make little sense biochemically, but can all be explained by a strong preference for amino acids with CG-rich codons in human. (Underlying diagram is after BETTS and RUSSELL 2003.)

DISCUSSION

To be viable as an explanation for isochores, a theory must be consistent with all of the following observations, from this work and the previous literature:

- i. Role of genes: Gene locations are special in isochores. Genes, both AT rich and CG rich, have more extreme compositions than their immediate intergenic flanking regions. A theory must explain Figure 6.
- ii. Composition asymmetry: CG isochores contain many AT-rich genes, while AT isochores contain few CG-rich genes. Not unrelated, CG isochores have a larger compositional variance on all scales than do AT isochores.
- iii. Gene functional correlations: Genes in CG isochores show a correlation between AT richness and GO function. Genes in AT isochores do not show such a correlation. On average, however, genes in the two isochores appear to have the same mixture of GO functions.
- iv. Spatial broken symmetry: How was any specific large region selected to become a CG isochore or not so selected?
- v. Evolutionary distance: There was not much more molecular evolution in the phylogenetic tree of iCG/CG genes than there was in iAT/AT genes. Rather, the branch leading to iCG/CG genes shows a strong mutational bias, not seen for iAT/AT genes. In the context of Kimura's theory of neutral evolution, this argues against positive selection pressure.

Properties i and v, both seemingly strongly supported by the data, seem contradictory. If the evolution of isochores is entirely neutral, then why are the genes in special locations? This question would be answered by a hypothetical (germ-line) mutational mechanism or repair process that acts differently in the vicinity of a gene than it does in a typical intergenic region: Either mutation rates near genes are higher or else they are (in CG isochores) more biased toward C and G.

Properties ii and iii, on the other hand, seem quite explainable. We start with a genome in a natural state of relative AT richness, as is seen in almost all animals except warm-blooded vertebrates. We suppose that a population of genes, about half, depend critically on regulatory or other functional mechanisms that depend on this "universal" AT richness. An example of such a mechanism may be regulation by micro-RNAs (ROBINS and PRESS 2005). Genes that do, or do not, require AT richness are randomly distributed on the genome. Something now happens, as posited by property iv: A dramatic mutational bias toward CG occurs in large genomic regions. Genes that are not dependent on AT-rich machinery become CG rich, that is, become iCG/CG genes. Those that are dependent experience purifying (negative) selection and remain AT rich, that is, become iCG/AT genes.

Property iv, requiring an explanation of what originally "painted" the propensity toward CG mutations onto large, coherent parts of the ancestral genome, is thus seen to be the fulcrum on which any explanation of isochores may tip.

Sequence data used in this article, including alignments, are available at <http://www.nr.com/bio/IsochoreSuppMat.html>.

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